

# Estimation techniques for MEV models with sampling of alternatives

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European Transport Conference

Glasgow, October 11-13, 2010



# Motivation

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- Discrete choice models with large or unknown choice sets require sampling of alternatives
- Consistent estimation is possible for MNL (McFadden, 1978)
- Sampling in non-logit models: can't be directly extended from MNL case
- Asymptotically unbiased estimator for nested logit proposed by Guevara and Ben-Akiva (2010)
- Bias can be reduced using bootstrapping techniques and importance sampling

# Outline

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1. Sampling of alternatives for MNL
2. MEV models
3. Sampling of alternatives for MEV models
4. Proposed techniques for reduced-bias estimation
  - 4.1 Bootstrapping
  - 4.2 Importance Sampling
5. Conclusions

# Sampling of alternatives for MNL

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- Choice probability with full choice set:

$$P(i) = \frac{e^{V_{ni}}}{\sum_{j \in C_n} e^{V_{nj}}}$$

- Choice probability with a sample  $D_n$  (McFadden, 1978):

$$P(i|D_n) = \frac{e^{\mu V_{ni} + \ln \pi(D_n|i)}}{\sum_{j \in D_n} e^{\mu V_{nj} + \ln \pi(D_n|j)}}$$

- Extension of this results for non-Logit models is not straightforward

# MEV models

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- Generating function  $G(e^{V_1}, \dots, e^{V_J})$
- Choice probability:

$$P_n(i) = \frac{e^{V_{in} + \ln G_i}}{\sum_{j \in C_n} e^{V_{jn} + \ln G_j}}$$

- where  $G_i = \frac{\partial G(e^{V_{1n}}, e^{V_{2n}}, \dots, e^{V_{J_n n}})}{\partial e^{V_{in}}}$

# MEV models

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- Different  $G$  functions generate different models:

- MNL: 
$$G = \sum_{j \in C_n} e^{\mu V_{jn}}$$

- Nested Logit: 
$$G = \sum_{m=1}^M \left( \sum_{j \in C_{mn}} e^{\mu_m V_{jn}} \right)^{\frac{\mu}{\mu_m}}$$

- Cross-nested Logit: 
$$G = \sum_{m=1}^M \left( \sum_{j \in C_{mn}} (\alpha_{jm} e^{V_{in}})^{\mu_m} \right)^{\frac{\mu}{\mu_m}}$$

# Sampling of alternatives for MEV models

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- Choice probability when considering a sample  $D_n$  (Bierlaire, Bolduc and McFadden, 2008):

$$P_n(i|D_n) = \frac{e^{V_{in} + \ln G_i + \ln \pi(D_n|i)}}{\sum_{j \in D_n} e^{V_{jn} + \ln G_j + \ln \pi(D_n|j)}}$$

- In many cases (NL, CNL)  $\ln G_i$  still depends on the full choice set  $C_n$

# Sampling of alternatives for MEV models

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- in the Nested Logit Case:

$$\ln G_{in} = \left( \frac{\mu}{\mu_{m(i)}} - 1 \right) \left( \ln \sum_{j \in C_{m(i)n}} e^{\mu_{m(i)} V_{jn}} \right) + \ln \mu + (\mu_{m(i)} - 1) V_{in}$$

- Logsum approximation (Guevara and Ben-Akiva, 2010):

$$\left( \ln \sum_{j \in C_{m(i)n}} e^{\mu_{m(i)} V_{jn}} \right) \approx \left( \ln \sum_{j \in D_{m(i)n}} w_{jn} e^{\mu_{m(i)} V_{jn}} \right)$$

- with  $w_{jn} = \frac{\tilde{n}_{jn}}{E_n(j)}$

# Sampling of alternatives for MEV models

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- $D_{mn}$  includes the chosen alternative and randomly sampled (without replacement) elements of the nest  $m$
- Estimation through maximum log-likelihood with the following choice probability

$$P_n(i|D_n) = \frac{e^{V_{in} + \ln G_i(D_{m(i)n}) + \ln \frac{|C_{m(i)}|}{|D_{m(i)n}|}}}{\sum_{j \in D_n} e^{V_{jn} + \ln G_j(D_{m(j)n}) + \ln \frac{|C_{m(j)}|}{|D_{m(j)n}|}}}$$

- where  $\ln G_i(D_{m(i)n})$  considers the approximated logsum

# Sampling of alternatives for MEV models

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- The approximated logsum generates asymptotically unbiased estimates  
→ biased parameters when the sample size is small (even when the true choice probabilities are used to calculate  $w_{jn}$ )
- Possible improvements for the approximated logsum:
  - Correction of the bias using Bootstrapping
  - Importance sampling of the elements in the logsum

# Bootstrapping

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- Simulation based technique for statistical inference of the properties of an estimator, from a sub-sample of observations
- Application to the approximated logsum:
  1. Estimation using the approximated logsum
  2. Re-sampling (with replacement) from the original sample of alternatives
  3. Re-calculation of the logsum with the new sample
  4. Calculation of the bias
  5. Re-estimation correcting for the bias

# Bootstrapping

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- Bootstrap estimator of the bias:

$$\rho_{mn} = \frac{1}{B} \sum_b^B \left( \ln \sum_{j \in D_{mn}^b} w_{jn} e^{\mu_m^0 V_{jn}(\beta^0)} \right) - \left( \ln \sum_{j \in D_{mn}} w_{jn} e^{\mu_m^0 V_{jn}(\beta^0)} \right)$$

- where
  - $\beta^0, \mu^0$  : set of parameters from the original estimation
  - $D_{mn}^b$  is the set of alternatives in each re-sampling instance ( $b$ )
  - $B$  is the number of re-sampling instances

# Bootstrapping

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- Estimation through maximum log-likelihood with the following choice probability

$$P_n(i|D_n) = \frac{e^{V_{in} + \ln \hat{G}_i(D_{m(i)n}) + \ln \frac{|C_{m(i)}|}{|D_{m(i)n}|}}}{\sum_{j \in D_n} e^{V_{jn} + \ln \hat{G}_j(D_{m(j)n}) + \ln \frac{|C_{m(j)}|}{|D_{m(j)n}|}}}$$

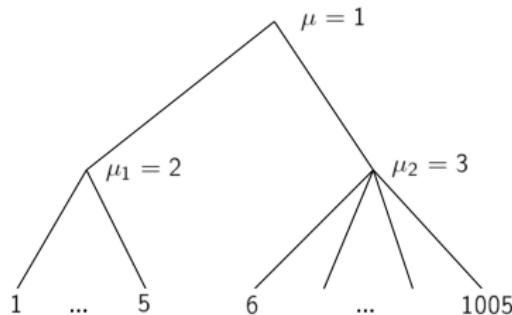
where:

$$\ln \hat{G}_i(D_{m(i)n}) = \left( \frac{\mu}{\mu_{m(i)}} - 1 \right) \left( \left( \ln \sum_{j \in D_{m(i)n}} w_{jn} e^{\mu_{m(i)} V_{jn}} \right) - \rho_{m(i)n} \right) + \ln \mu + (\mu_{m(i)} - 1) V_{in}$$

# Bootstrapping: Experiment

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- Nested logit:



- Utility:  $V_{in} = \beta_a a_{in} + \beta_b b_{in}$
- Attributes:  $a_{in}, b_{in} \sim U(-1, 1)$
- True parameters  $\beta_a = 1, \beta_b = 1, \mu_1 = 2, \mu_2 = 3$
- Sampling of alternatives within nest 2

# Bootstrapping: Results

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- Results from Monte Carlo experiment using approximated logsum (sample size = 5):

parameter	average value	std-error	true value	t-test
$\beta_a$	0.855	0.082	1	1.773
$\beta_b$	0.843	0.068	1	2.288 *
$\mu_1$	2.569	0.581	2	0.978
$\mu_2$	3.622	0.272	3	2.290 *

\* Biased estimates

# Bootstrapping: Results

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- Results after bootstrapping (sample size = 5):

parameter	average value	std-error	true value	t-test
$\beta_a$	0.953	0.079	1	0.595
$\beta_b$	0.957	0.079	1	0.548
$\mu_1$	2.264	0.517	2	0.511
$\mu_2$	3.224	0.229	3	0.974

- significant reduction of the bias

# Importance sampling

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- The bias can be reduced with a better sample for the approximation of the logsum
- The sample of alternatives in the logsum does not have to be the same as the sample of alternatives for the choice set
- Method:
  1. Random sampling of alternatives for the elements in the logsum
  2. Estimation using approximated logsum  $\rightarrow \beta^0, \mu^0$
  3. Importance sampling of alternatives for the logsum following  $P(\beta^0, \mu^0)$
  4. Re-estimation

# Importance sampling

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- First estimation:

$$P_n(i|D_n) = \frac{e^{V_{in} + \ln G_i(L_{m(i)n}) + \ln \frac{|c_{m(i)}|}{|D_{m(i)n}|}}}{\sum_{j \in D_n} e^{V_{jn} + \ln G_j(L_{m(j)n}) + \ln \frac{|c_{m(j)}|}{|D_{m(j)n}|}}}$$

- where  $L_{m(i)n}$  is randomly generated sample of alternatives for the logsum ( $|L_{m(i)n}| = |D_{m(i)n}|$ )
- From this first estimation we get  $\beta^0, \mu^0$

# Importance sampling

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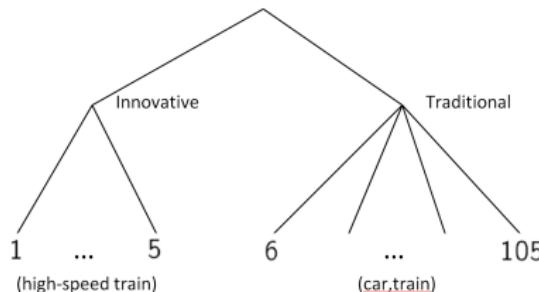
- The sampling for the logsum is done again, following a MNL inside the nest using the previously estimated parameters  $(\beta^0, \mu^0)$

$$g_n(i|m) = \frac{e^{V_{ni}(\beta^0, \mu^0)}}{\sum_{j \in C_m} e^{V_{nj}(\beta^0, \mu^0)}}$$

- The elements in the sample for the choice set remain the same
- A re-estimation is performed

# Importance sampling: Experiment

- Synthetic data generated from a survey to evaluate a high speed train in Switzerland



- $V_{hs} = \beta_{cost} C_{hs} + \beta_{time\_T} TT_{hs} + \beta_{headway} HE_{hs}$
- $V_{car} = \beta_{cost} C_{car} + \beta_{time\_c} TT_{car}$
- $V_{train} = \beta_{cost} C_{train} + \beta_{time\_T} TT_{train} + \beta_{headway} HE_{train}$

# Importance sampling: Results

- Results from Monte Carlo experiment using approximated logsum (sample size = 5):

parameter	average value	std-error	true value	t-test
$\beta_{cost}$	-1.253	0.152	-0.849	2.666 *
$\beta_{time\_C}$	-2.958	0.359	-1.760	3.388 *
$\beta_{time\_T}$	-2.708	0.306	-1.840	2.835 *
$\beta_{headway}$	-0.967	0.217	-0.496	2.165 *
$\mu_1$ (innovative)	1.220	0.160	2	4.873 *
$\mu_2$ (traditional)	3.146	0.368	4	2.318 *

\* Biased estimates

# Importance sampling: Results

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- Results from Monte Carlo experiment using importance sampling (sample size = 5):

parameter	average value	std-error	true value	t-test
$\beta_{cost}$	-0.930	0.135	-0.849	0.560
$\beta_{time\_C}$	-1.997	0.321	-1.760	0.736
$\beta_{time\_T}$	-2.008	0.314	-1.840	0.535
$\beta_{headway}$	-0.592	0.143	-0.496	0.672
$\mu_1$ (innovative)	1.766	0.359	2	0.652
$\mu_2$ (traditional)	3.503	0.430	4	1.155

- Significant reduction of the bias

# Conclusions

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- Two methods that reduce the bias for MEV model estimation were presented
- Bootstrapping reduces the bias of the approximated logsum
  - bootstrapped results will depend on the quality of the original estimator
- Importance sampling of the elements in the logsum allows to find unbiased estimates
  - different sample for the choice set and the elements in the logsum
- Further work:
  - Test other correlation structures (e.g. Cross-nestedlogit)
  - Estimation over real data

Thank you