

# A Unimodal Ordered Logit model for ranked choices

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# Outline

- Introduction
- Background
- Methodology: Unimodal logit
- Case study: Crash severity model
- Conclusion

# Introduction

- Ordinal scale responses capture qualitative user feedback
- Responses have inherent correlation between alternatives [Small, 1987]



## Examples

PT satisfaction, driver star-rating (ride-hailing), crash severity...

- [Krueger et al., 2019, Tirachini and del Río, 2019, Fu, 2020, Loa and Habib, 2021]

# Background

[McCullagh, 1980]

Proportional odds model

- Contiguous intervals on a continuous scale
- Points of division assumed to be unknown



[Small, 1987]

Ordered logit, Generalized ordered logit

- Define a latent variable ( $y^*$ ) that varies across the contiguous intervals
- $y^* \leftarrow$  exogenous features of the response
- $y^* = \sum_m \beta_m X_m$
- Choice prob. = probability of lying in any of the intervals

## Modelling non-ordered choices

Assume that there are  $J$  alternatives ( $i = 1, \dots, J$ )

- Denote  $y_{ni} = 1$  if individual  $n$  is ranked in  $i$  and  $y_{ni} = 0$  otherwise
- $n = 1, \dots, N$ ,  $U_{n1}, \dots, U_{nJ}$ ,  $U_{ni} \geq \max\{U_{n1}, \dots, U_{nJ}\}$
- $U_{ni} = V_{ni} + \varepsilon_{ni}$ ,  $\varepsilon_{ni} \sim \text{Gumbel}(0, 1)$  i.i.d.

### Multinomial logit model

$$P(y_{ni} = 1) = \frac{\exp(V_{ni})}{\sum_{j=1}^J \exp(V_{nj})}$$

For choices with natural ordering, i.i.d. assumption does not hold

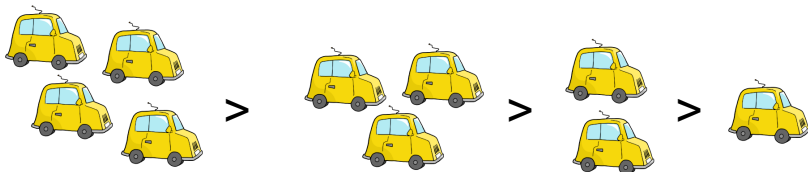
- Standard MNL model is not suitable in this context of ranked choices

# Modelling ordered choices

## Example

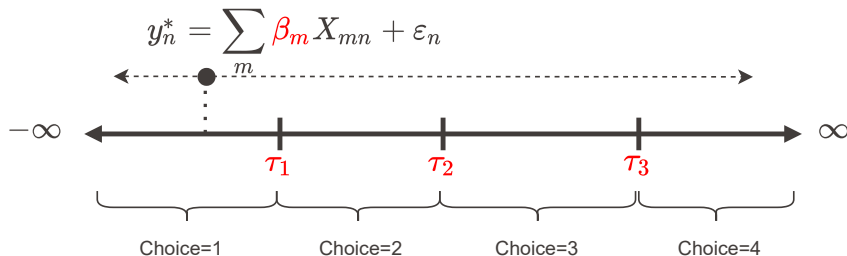
Vehicle ownership model [Sheffi, 1979]

- $j$  is the number of household vehicles ( $j = 1, 2, 3, \dots$ )
- $n$  would prefer  $j$  over  $j - 1$ ,  $j - 1$  over  $j - 2$ , and so on..
- Correlation between choices not captured (MNL)



# Modelling ordered choices

## Ordered logit



## Estimating thresholds

$$\tau_1 = 0, \tau_2 = \tau_1 + \Delta_2, \tau_3 = \tau_2 + \Delta_3$$

# Modelling ordered choices

## Ordered logit

- The difference between thresholds (e.g. between  $\tau_2$  and  $\tau_3$ ) are assumed to be the same for all respondents
- Parameters  $\beta_m$  are constant across all respondents
- Typically set threshold  $\tau_1 = 0$  for model identification

## Generalized ordered logit [Eluru et al., 2008]

- latent variable combines alt. specific and generic parameters

$$y_n^* = \sum_m \beta_m X_{mn} + \sum_m \beta_{im} X_{mn} + \varepsilon_n$$

- Thresholds are functions of exogenous variables:

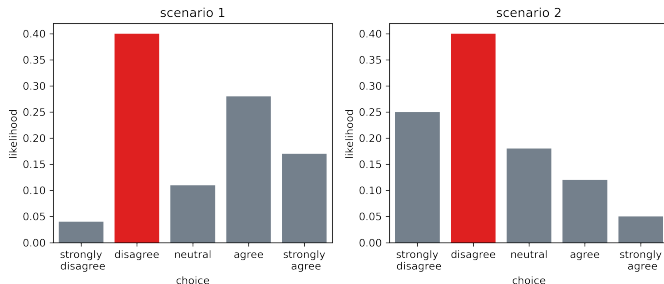
$$\tau_i = \tau_{i-1} + \exp\left(\sum_m \delta_{im} Z_{imn}\right)$$



## Other models for ordered choices

- Generalized Extreme Value (GEV) [McFadden, 1977]
- Ordered GEV [Small, 1987]
- Dogit model [Gaundry and Dagenais, 1979]
- Dogit OGEV model [Fry and Harris, 2005]

# A different approach for ordered choices



## Maximum likelihood estimation

$$\ln(P(y_{\text{scenario 1}} = 2)) = \ln(P(y_{\text{scenario 2}} = 2))$$

- Both result in identical max likelihood, but probability mass function (pmf) is different

# Unimodality in ordered choices

## Properties

- Unimodality: A single highest value
- Specifically, the *a posteriori* choice probabilities are unimodal

Natural ordering of choices is captured in the model if there exist an integer  $c \in J$  such that:

- $p(y_{ni}|X) \geq p(y_{ni+1}|X)$ , for all  $i \geq c$  and,
- $p(y_{ni-1}|X) \leq p(y_{ni}|X)$ , for all  $i \leq c$

# Unimodality in ordered choices

## Poisson pmf

The probability of  $i$  occurrences of an event in a set of  $N$  observations is defined as:

$$P(i) = \frac{\lambda^i \exp(-\lambda)}{i!}, \text{ for } i = 0, 1, 2, ..$$

# Unimodal logit

Applying a unimodal constraint in the utility function:

$$\begin{aligned}
 U_{in} &= V_{in} + \ln(P(i)) + \varepsilon_{in} \\
 &= V_{in} + \ln\left(\frac{\lambda^i \exp(-\lambda)}{i!}\right) + \varepsilon_{in} \\
 &= V_{in} + \underbrace{i \ln(\lambda) - \lambda - \ln(i!)}_{\text{error component } f(\lambda, i)} + \varepsilon_{in}
 \end{aligned}$$

ec: capture correlations among utilities of alternatives

## Conditions

- $\lambda$  is positive
- $\lambda = f(y_n^*) = \ln(1 + \exp(y_n^*))$

# Unimodal logit

Expressed as a MNL choice probability:

$$P(y_{ni} = 1) = \frac{\exp(\mu\Phi_{in})}{\sum_{j=1}^J \exp(\mu\Phi_{jn})}$$

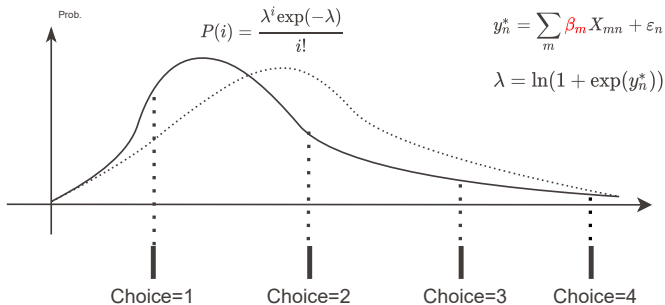
$$\Phi_{in} = V_{in} + i \ln(\lambda) - \lambda - \ln(i!) + \beta_{i0}$$

## Behavioural interpretation

Utilities of alternatives are corrected for proximity from the selected choice  $i$

# Unimodal logit

## Unimodal logit



## Zero-truncated Poission (ZTP) pmf

When a choice set has a “zero” option

- Example: Number of items in a shopping cart include a “no purchase” option

A ZTP Unimodal logit has the following pmf:

$$P(i|i > 0) = \frac{\lambda^i \exp(-\lambda)}{i!(1 - \exp(-\lambda))}, \text{ for } i = 1, 2, 3, \dots$$

$$U_{in} = V_{in} + i \ln(\lambda) - \lambda - \ln(i!) - \ln(1 - \exp(-\lambda)) + \varepsilon_{in}$$



# Crash severity model

[City of Tempe, 2018]

Open dataset: High Severity Traffic Crash Data Report

- 39,793 records (2012–2019)
- Five severity levels
  - 1: No injury, 2: possible injury, 3: minor injury, 4: major injury, 5: fatal
- 28 crash and environmental features used (after data cleaning)

## Models

Estimation using Biogeme [Bierlaire, 2020]

- Ordered logit
- Unimodal logit
- Zero truncated unimodal logit

# Crash severity model

## Model Evaluation

### Goodness-of-fit

- Pseudo R-squared measure ( $\rho^2$ )

$$\rho^2 = 1 - \frac{\ln LL(\hat{\beta})}{\ln LL(\bar{\beta})}$$

- Bayesian Information Criterion (BIC)

$$BIC = -2LL(\beta) + M \ln(Q)$$

### Out-of-sample accuracy

- Discrete classification accuracy
- Geometric mean probability of correct assignment (GMPCA) [Hillel, 2019]
- Quadratic Weighted Kappa (QWK) [Cohen, 1968]

# Model results

## Abridged results (1)

Variables	Ordered Logit		Unimodal		Zero-trunc Unimodal	
	values	rob_tTest	values	rob_tTest	values	rob_tTest
age	-0.008	-10.184	-0.015	-22.834	-0.017	-18.351
alcohol	0.384	5.02	0.379	4.625	0.524	4.918
<b>cause_distraction</b>	<b>0.08</b>	<b>1.013</b>	<b>-0.287</b>	<b>-4.749</b>	-0.249	<b>-2.83</b>
<b>cause_speeding</b>	-0.027	<b>-0.543</b>	-0.271	<b>-6.78</b>	-0.28	<b>-4.824</b>
cause_turn	-0.153	-1.832	-0.355	-6.05	-0.411	-4.611
cause_yield	-0.108	-2.038	-0.341	-8.166	-0.4	-6.784
<b>type_cyclist</b>	1.46	<b>17.722</b>	0.619	<b>5.289</b>	0.804	<b>6.265</b>
<b>type_driverless</b>	-0.52	<b>-1.465</b>	-1.478	<b>-7.631</b>	-1.744	<b>-5.267</b>
type_pedestrian	1.596	7.657	3.838	8.122	3.066	6.065

# Model results

## Abridged results (2)

Variables	Ordered Logit		Unimodal		Zero-trunc Unimodal	
	values	rob_tTest	values	rob_tTest	values	rob_tTest
ASC_noinjury (1)			ref.		ref.	
ASC_possinjury (2)			3.673	93.809	2.452	68.248
ASC_nonincap (3)			4.117	104.798	3.446	97.33
ASC_incap (4)			2.449	43.533	1.907	35.544
ASC_fatal (5)			0.788	7.193	0.319	2.889
tau1	0.0	0.0				
delta2	2.611	68.111				
delta3	3.31	39.596				
delta4	2.303	14.98				

# Model results

	Ordered Logit	Unimodal	Zero-trunc Unimodal
Log likelihood	-17148.44	<b>-13471.31</b>	-16731.04
BIC	34628.6	<b>27274.4</b>	33793.8
$\rho^2$	0.665	<b>0.737</b>	0.673
Optimization time	0:01:02.27	0:06:26.2	0:07:40.4
Discrete Class. Acc.	0.839	0.842	0.826
GMPCA	0.581	<b>0.653</b>	0.59
QWK	0.758	<b>0.805</b>	0.787
20% out-of-sample data used			

# Conclusion

We introduce a new form of choice model for ordered choices

- Unimodal constraint on the *a posteriori* distribution
- Similar  $\beta$  interpretations as Ordered logit

## Case study

- Able to capture the influence of relevant crash severity characteristics: driving speed, distracted driving and driverless vehicles
- Exhibit better model fit and forecasting accuracy

## Future work

- Negative binomial distribution
- Combination with other error correction functions

Thank you for your attention

Estimated models, cleaned data and data analysis are available at:  
<https://github.com/mwong009/unimodal-logit>

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