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# The estimation of generalized extreme value models from choice-based samples

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# Outline

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- Introduction
- Sampling
- Estimation
- Multivariate (generalized) extreme value models
- Illustrations

# Introduction

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- Sampling is never random in practice
- Choice-based samples are convenient in transportation analysis
- Estimation is an issue
- Main references:
  - Manski and Lerman (1977)
  - Manski and McFadden (1981)
  - Cosslett (1981)
  - Ben-Akiva and Lerman (1985)

# Sampling: context

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- Discrete choice model,  $J$  alternatives
- Independent variables:  $x$
- Dependent variable (choice):  $i$
- Model:

$$\Pr(i|x, \theta) = P(i|x, \theta)$$

- Unknown parameters:  $\theta$
- Joint distribution of  $(i, x)$  in the population

$$\Pr(i, x|\theta) = P(i|x, \theta)p(x).$$

# Sampling: stratification

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- Population partitioned into  $G$  groups
- Individuals randomly selected within each group
- Population size:  $N_P$
- % of ind. from group  $g$  in population:  $W_g$
- Sample size:  $N_s$
- % of ind. from group  $g$  in sample:  $H_g$
- Probability to be in the sample:  $r_g = \frac{H_g N_s}{W_g N_P}$ .

# Sampling strategies

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## SRS Simple random sampling

- Only one group.
- $H_g = W_g$ ,
- $r_g = r = N_s/N_P$ .

## xss Exogenous stratified sampling

- Groups characterized by  $x$
- $W_g = \int_{x \in X_g} p(x) dx$
- $r_g$  does not depend on  $\theta$ .

# Sampling strategies

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## ESS Endogenous stratified sampling

- Groups characterized by  $i$
- $W_g$  does not simplify
- $r_g$  depends on  $\theta$

## XESS Exogenous and endogenous stratified sampling

- Groups characterized both by  $x$  and  $i$
- $W_g$  does not simplify
- $r_g$  depends on  $\theta$

# Sampling of alternatives

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- Analyze choice as if limited to  $\mathcal{B} \subseteq \mathcal{C}$
- $\mathcal{B}$  is drawn with prob.  $\pi(\mathcal{B}|i, x)$
- Positive conditioning property:

$$\pi(\mathcal{B}|i, x) > 0 \Rightarrow \pi(\mathcal{B}|j, x) > 0 \quad \forall j \in \mathcal{B}.$$

- Appropriate sampling:

$$\pi(\mathcal{B}|i, x) > 0 \Rightarrow r_{g(i,x)} > 0$$



# Sampling

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Probability that a population member with configuration  $(i, x)$  is sampled, and is assigned the truncated choice set  $\mathcal{B}$ :

$$R(i, x, \mathcal{B}, \theta) = \Pr(s, \mathcal{B} | i, x, \theta) = r_{g(i,x)}(\theta) \pi(\mathcal{B} | i, x).$$

# Estimation

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## Conditional Maximum Likelihood (CML) Estimator

$$\begin{aligned}\max_{\theta} \mathcal{L}(\theta) &= \sum_{n=1}^N \ln \Pr(i_n | x_n, \mathcal{B}_n, s, \theta) \\ &= \sum_{n=1}^N \ln \frac{R(i_n, x_n, \mathcal{B}_n, \theta) P(i_n | x_n, \theta)}{\sum_{j \in \mathcal{B}_n} R(j, x_n, \mathcal{B}_n, \theta) P(j | x_n, \theta)}\end{aligned}$$

In practice,  $R(i_n, x_n, \mathcal{B}_n, \theta)$  cannot be computed, namely because it requires  $p(x)$

# Estimation

Assume that  $R(i, x, \mathcal{B}, \theta)$  can be written as

$$R(i, x, \mathcal{B}, \theta) = Q(i, x, \mathcal{B})S(i, x, \mathcal{B}, \theta).$$

Pseudo-likelihood function

$$\hat{\mathcal{L}} = \sum_{n=1}^N Q(i_n, x_n, \mathcal{B}_n)^{-1} \ln \frac{S(i_n, x_n, \mathcal{B}_n, \theta) P(i_n | x_n, \theta)}{\sum_{j \in \mathcal{B}_n} S(j, x_n, \mathcal{B}_n, \theta) P(j | x_n, \theta)}$$

- $Q = 1$ : CML by Manski & McFadden (1981)
- $S = 1$ : WESML by Manski & Lerman (1977)

# Estimation of MEV models

- Let  $G$  be the generating function of a MEV model
- Let

$$G_i(x, \beta, \gamma) = \frac{\partial G}{\partial e^{V_i(x, \beta)}} \left( e^{V_1(x, \beta)}, \dots, e^{V_J(x, \beta)}; \gamma \right).$$

- The main term in the CML formulation is:

$$\frac{S(i, x, \mathcal{B}, \theta) P(i|x, \theta)}{\sum_{j \in \mathcal{B}} S(j, x, \mathcal{B}, \theta) P(j|x, \theta)} = \frac{e^{V_i(\beta) + \ln G_i(x, \beta, \gamma) + \ln S(i, x, \mathcal{B}, \theta)}}{\sum_{j \in \mathcal{B}} e^{V_j(\beta) + \ln G_j(x, \beta, \gamma) + \ln S(j, x, \mathcal{B}, \theta)}}$$

# Estimation of MEV models

- The term needed for CML is MNL-like
- Case of MNL model:  $G_i = 0$ .

$$\frac{S(i, x, \mathcal{B}, \theta) P(i|x, \theta)}{\sum_{j \in \mathcal{B}} S(j, x, \mathcal{B}, \theta) P(j|x, \theta)} = \frac{e^{V_i(\beta) + \ln S(i, x, \mathcal{B}, \theta)}}{\sum_{j \in \mathcal{B}} e^{V_j(\beta) + \ln S(j, x, \mathcal{B}, \theta)}}.$$

- Well-known result: if ESML is used, only constants are biased
- Question: does this generalize to all MEV?
- Answer: not quite...

# Estimation of MEV models

- The  $V$ 's are shifted in the main formula

$$\frac{e^{V_i(\beta) + \ln G_i(x, \beta, \gamma) + \ln S(i, x, \mathcal{B}, \theta)}}{\sum_{j \in \mathcal{B}} e^{V_j(\beta) + \ln G_j(x, \beta, \gamma) + \ln S(j, x, \mathcal{B}, \theta)}}.$$

- ... but not in the  $G_i$

$$G_i(x, \beta, \gamma) = \frac{\partial G}{\partial e^{V_i(x, \beta)}} \left( e^{V_1(x, \beta)}, \dots, e^{V_J(x, \beta)}; \gamma \right).$$

- **ESML will not produce consistent estimates on non-MNL MEV models.**

# Estimation of MEV models

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$$\frac{e^{V_i(\beta) + \ln G_i(x, \beta, \gamma) + \ln S(i, x, \mathcal{B}, \theta)}}{\sum_{j \in \mathcal{B}} e^{V_j(\beta) + \ln G_j(x, \beta, \gamma) + \ln S(j, x, \mathcal{B}, \theta)}}$$

- New idea: estimate  $\ln S(i, x, \mathcal{B}, \theta)$  from data
- Cannot be done with classical software
- But easy to implement due to the MNL-like form

# Illustration

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- Pseudo-synthetic data
- Data base: SP mode choice for future highspeed train in Switzerland (Swissmetro)
- Alternatives:
  1. Regular train (TRAIN),
  2. Swissmetro (SM), the future high speed train,
  3. Driving a car (CAR).
- Generation of a synthetic population of 507600 individuals



# Illustration

- Attributes are random perturbations of actual attributes
- Assumed true choice model: NL

Param.	Value	Alternatives		
		TRAIN	SM	CAR
ASC_CAR	-0.1880	0	0	1
ASC_SM	0.1470	0	1	0
B_TRAIN_TIME	-0.0107	travel time	0	0
B_SM_TIME	-0.0081	0	travel time	0
B_CAR_TIME	-0.0071	0	0	travel time
B_COST	-0.0083	travel cost	travel cost	travel cost

# Illustration

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- Nesting structure:

	$\mu_m$	TRAIN	SM	CAR
NESTA	2.27	1	0	1
NESTB	1.0	0	1	0

# Illustration

- 100 samples drawn from the population

Strata	$W_g N_P$	$W_g$	$H_g$	$H_g N_s$	$R_g$
TRAIN	67938	13.4%	60%	3000	4.42E-02
SM	306279	60.3%	20%	1000	3.26E-03
CAR	133383	26.3%	20%	1000	7.50E-03
Total	507600	1	1	5000	

- Estimation of 100 models
- Empirical mean and std dev of the estimates

# Illustration

	True	ESML			New estimator		
		Mean	<i>t</i> -test	Std. dev.	Mean	<i>t</i> -test	Std. dev.
ASC_SM	0.1470	-2.2479	-25.4771	0.0940	-2.4900	-23.9809	0.1100
ASC_CAR	-0.1880	-0.8328	-7.3876	0.0873	-0.1676	0.1581	0.1292
BCOST	-0.0083	-0.0066	2.6470	0.0007	-0.0083	0.0638	0.0008
BTIME_TRAIN	-0.0107	-0.0094	1.4290	0.0009	-0.0109	-0.1774	0.0009
BTIME_SM	-0.0081	-0.0042	3.1046	0.0013	-0.0080	0.0446	0.0014
BTIME_CAR	-0.0071	-0.0065	0.9895	0.0007	-0.0074	-0.3255	0.0007
NestParam	2.2700	2.7432	1.7665	0.2679	2.2576	-0.0609	0.2043
S_SM_Shifted	-2.6045						
S_CAR_Shifted	-1.7732				-1.7877	-0.0546	0.2651
ASC_SM+S_SM	-2.4575				-2.4900	-0.2958	0.1100

# Conclusions

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- Except in very specific cases, ESML provides biased estimated for non-MNL MEV models
- Due to the MNL-like form of the MEV model, a new simple estimator has been proposed
- It allows to estimate selection bias from the data