### Variance reduction

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### **Example**

Use simulation to compute

$$I = \int_0^1 e^x \ dx$$

- We know the solution: e 1 = 1.7183
- Simulation: consider draws two by two
- Let  $r_1, \ldots, r_R$  be independent draws from U(0, 1).
- Let  $s_1, \ldots, s_R$  be independent draws from U(0, 1).

$$I \approx \frac{1}{R} \sum_{i=1}^{R} \frac{e^{r_i} + e^{s_i}}{2}$$

- Use R = 10'000 (that is, a total of 20'000 draws)
- Mean over R draws from  $(e^{r_i} + e^{s_i})/2$ : 1.720,variance: 0.123



### **Example**

- Now, use half the number of draws
- Idea: if  $X \sim U(0,1)$ , then  $(1-X) \sim U(0,1)$
- Let  $r_1, \ldots, r_R$  be independent draws from U(0, 1).

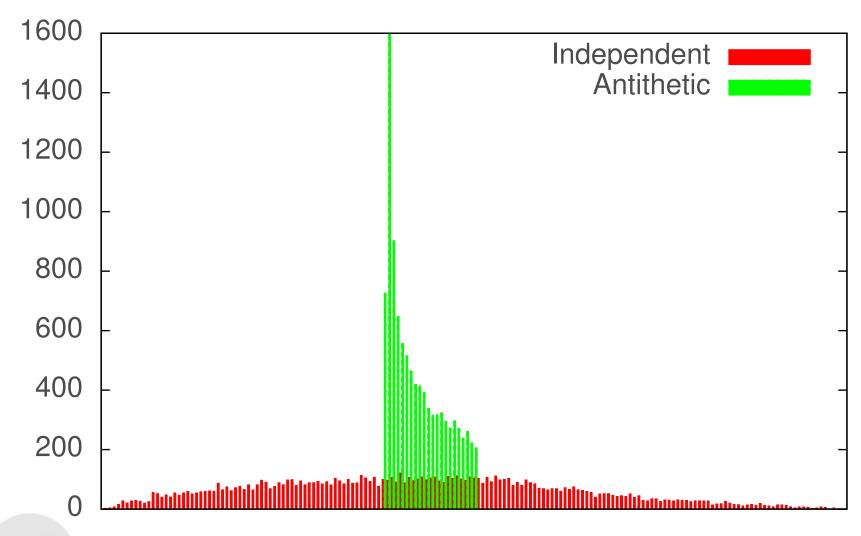
$$I \approx \frac{1}{R} \sum_{i=1}^{R} \frac{e^{r_i} + e^{1-r_i}}{2}$$

- Use R = 10'000
- Mean over R draws of  $(e^{r_i} + e^{1-r_i})/2$ : 1.7183,variance: 0.00388
- Compared to: mean of  $(e^{r_i} + e^{s_i})/2$ : 1.720, variance: 0.123





# **Example**





### **Antithetic draws**

- Let  $X_1$  and  $X_2$  i.i.d r.v. with mean  $\theta$
- Then

$$\operatorname{Var}\left(\frac{X_1 + X_2}{2}\right) = \frac{1}{4}\left(\operatorname{Var}(X_1) + \operatorname{Var}(X_2) + 2\operatorname{Cov}(X_1, X_2)\right)$$

- If  $X_1$  and  $X_2$  are independent, then  $Cov(X_1, X_2) = 0$ .
- If  $X_1$  and  $X_2$  are negatively correlated, then  $Cov(X_1, X_2) < 0$ , and the variance is reduced.



Independent draws

• 
$$X_1 = e^U$$
,  $X_2 = e^U$ 

$$Var(X_1) = Var(X_2) = E[e^{2U}] - E[e^{U}]^2$$

$$= \int_0^1 e^{2x} dx - (e-1)^2$$

$$= \frac{e^2 - 1}{2} - (e-1)^2$$

$$= 0.2420$$

$$Cov(X_1, X_2) = 0$$

$$\operatorname{Var}\left(\frac{X_1 + X_2}{2}\right) = \frac{1}{4}\left(0.2420 + 0.2420\right) = 0.1210$$





- Antithetic draws
- $X_1 = e^U$ ,  $X_2 = e^{1-U}$

$$Var(X_1) = Var(X_2) = 0.2420$$

$$Cov(X_1, X_2) = E[e^U e^{1-U}] - E[e^U]E[e^{1-U}]$$
  
=  $e - (e-1)(e-1)$   
=  $-0.2342$ 

$$\operatorname{Var}\left(\frac{X_1 + X_2}{2}\right) = \frac{1}{4}\left(0.2420 + 0.2420 - 2\ 0.2342\right) = 0.0039$$





### Antithetic draws: generalization

Suppose that

$$X_1 = h(U_1, \dots, U_m)$$

where  $U_1, \ldots U_m$  are i.i.d. U(0,1).

Define

$$X_2 = h(1 - U_1, \dots, 1 - U_m)$$

- X<sub>2</sub> has the same distribution as X<sub>1</sub>
- If h is monotonic in each of its coordinates, then  $X_1$  and  $X_2$  are negatively correlated.
- If h is not monotonic, there is no guarantee that the variance will be reduced.





### **Another example**

$$I = \int_0^1 \left( x - \frac{1}{2} \right)^2 dx$$

Antithetic draws:

$$X_1 = \left(U - \frac{1}{2}\right)^2, \ X_2 = \left((1 - U) - \frac{1}{2}\right)^2$$

• The covariance is positive:

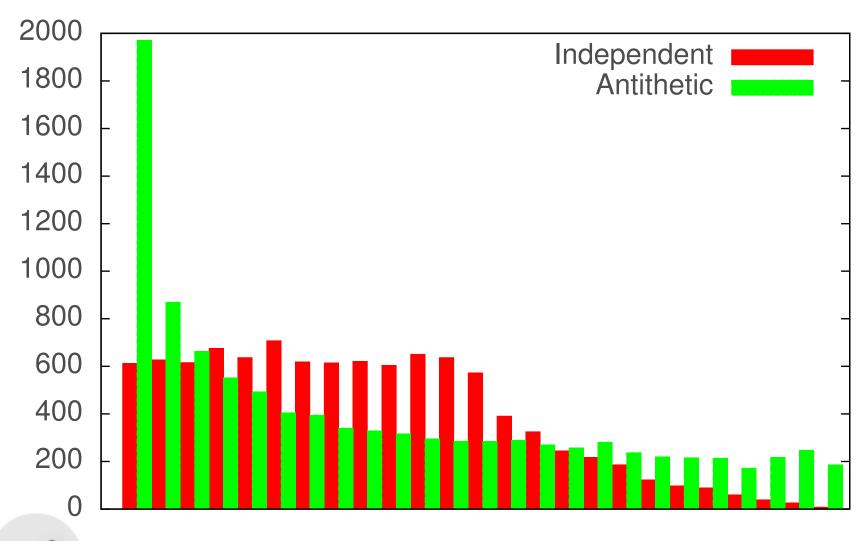
$$Cov(X_1, X_2) = \frac{1}{180} > 0.$$

• The variance will therefore be (slightly) increased!





### **Another example**



- We use simulation to estimate  $\theta = E[X]$ , where X is an output of the simulation
- Let Y be another output of the simulation, such that we know  $\mathrm{E}[Y] = \mu$
- We consider the quantity:

$$Z = X + c(Y - \mu).$$

- By construction, E[Z] = E[X]
- Its variance is

$$Var(Z) = Var(X + cY) = Var(X) + c^{2} Var(Y) + 2c Cov(X, Y)$$

• Find c such that Var(Z) is minimum



First derivative:

$$2c\operatorname{Var}(Y) + 2\operatorname{Cov}(X,Y)$$

Zero if

$$c^* = -\frac{\operatorname{Cov}(X, Y)}{\operatorname{Var}(Y)}$$

Second derivative:

$$2\operatorname{Var}(Y) > 0$$

We use

$$Z^* = X - \frac{\operatorname{Cov}(X, Y)}{\operatorname{Var}(Y)}(Y - \mu).$$

Its variance

$$\operatorname{Var}(Z^*) = \operatorname{Var}(X) - \frac{\operatorname{Cov}(X, Y)^2}{\operatorname{Var}(Y)} \le \operatorname{Var}(X)$$





#### In practice...

- Cov(X, Y) and Var(Y) are usually not known.
- We can use their sample estimates:

$$\widehat{\text{Cov}}(X,Y) = \frac{1}{n-1} \sum_{r=1}^{R} (X_r - \bar{X})(Y_r - \bar{Y})$$

and

$$\widehat{\text{Var}}(Y) = \frac{1}{n-1} \sum_{r=1}^{R} (Y_r - \bar{Y})^2.$$



#### In practice...

• Alternatively, use linear regression

$$X = aY + b + \varepsilon$$

where  $\varepsilon \sim N(0, \sigma^2)$ .

The least square estimators of a and b are

$$\hat{a} = \frac{\sum_{r=1}^{R} (X_r - \bar{X})(Y_r - \bar{Y})}{\sum_{r=1}^{R} (Y_r - \bar{Y})^2}$$

$$\hat{b} = \bar{X} - \hat{a}\bar{Y}.$$

Therefore

$$c^* = -\hat{a}$$





Moreover,

$$\hat{b} + \hat{a}\mu = \bar{X} - \hat{a}\bar{Y} + \hat{a}\mu 
= \bar{X} - \hat{a}(\bar{Y} - \mu) 
= \bar{X} + c^*(\bar{Y} - \mu) 
= \hat{\theta}$$

• Therefore, the control variate estimate  $\widehat{\theta}$  of  $\theta$  is obtained by the estimated linear model, evaluated at  $\mu$ .



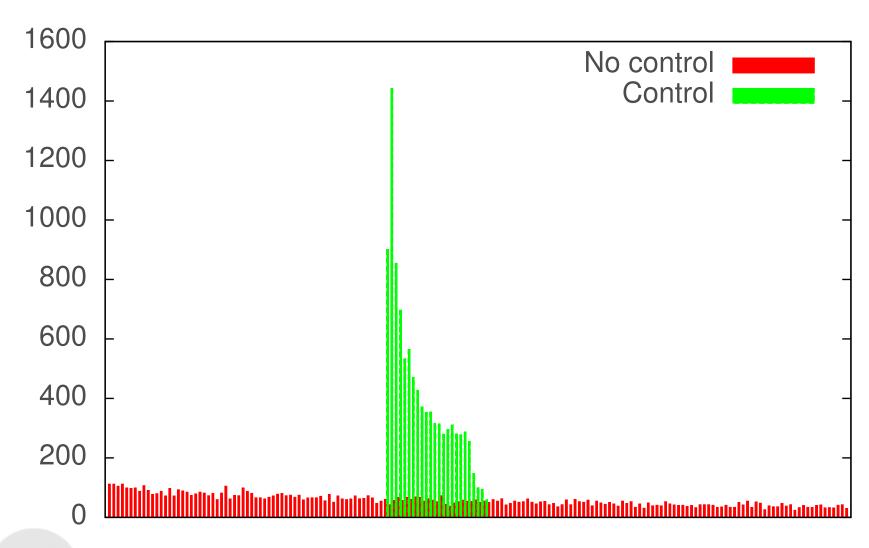


- Use simulation to compute  $I = \int_0^1 e^x \ dx$
- $\bullet X = e^U$
- Y = U, E[Y] = 1/2, Var(Y) = 1/12
- $Cov(X, Y) = (3 e)/2 \approx 0.14$
- Therefore, the best *c* is

$$c^* = -\frac{\text{Cov}(X, Y)}{\text{Var}(Y)} = -6(3 - e) \approx -1.69$$

- Test with R = 10'000
- Result of the regression:  $\hat{a} = 1.6893$ ,  $\hat{b} = 0.8734$
- Estimate:  $\hat{b} + \hat{a}/2 = 1.7180$ , Variance: 0.003847 (compared to 0.24)







# Variance reductions techniques

- Conditioning
- Stratified sampling
- Importance sampling
- Draw recycling

In general: correlation helps!



