

# Spatial modelling of origin-destination commuting flows in Switzerland

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## Abstract

### Objective and methodology

This paper presents a direct modelling approach for origin-destination public transportation commuting flows with possible endogenous regressors for the case of Switzerland. Its purpose is to improve the gravity modelling approach for OD flows by applying a spatial autoregressive regression model and testing different spatial weighting schemes. To the best of our knowledge, there has been no prior application of such advanced models in the context of transport demand modelling for public transport. Methodologically, in the first step a gravity model is developed and tested for the presence of spatial autocorrelation in its residuals. Subsequently, variants of a spatial lag model with different spatial weighting schemes are developed. Furthermore, we test a variable based on mean income differences on its ability to describe interregional demand patterns. In addition, we treat for the endogenous nature of the newly constructed variable. We are also testing its ability to serve as the basis for the construction of the spatial weight matrix, thus replacing the commonly used travel time / distance metric. On the modelling front, we use an Ordinary least squares (OLS) estimator for the gravity model, while a Generalized Method of Moments and Instrumental Variable (GMM/IV) (IV) estimator for the spatial models is employed in order to obtain unbiased and consistent parameter estimates. Last, we evaluate various models' goodness-of-fit measures and in-sample predictive accuracies by comparing among each other as well as to those of a state-of-the-practice transport model (as provided by national spatial planning bureau (NPVM)). This comparison can allow us to draw solid conclusion with respect to the suitability of the presented method for predicting commuting flows.

### Case study

In brief, a case study for public transport commuting flows in Switzerland is designed to illustrate the concept of OD flow modelling, based on travel-to-work trips data from the Federal Census of 2000. The data cover 2896 Swiss municipalities and contains over 250'000 observations on their initial form. However, it does not fill the whole flow matrix that contains  $2896^2 = 8,386,816$  flows. For the remaining OD pairs we assume zero-valued travel flows. An important aspect is the issue with how to deal with zero flows. A large fraction of zero-valued OD flows would definitely point towards a Poisson or a

59 (zero-inflated) negative Binomial interaction model. However, neither a Poisson nor a negative  
60 Binomial spatial autoregressive regression model for OD flows has been developed so far. We include  
61 income differences between Swiss communes as an explanatory variable in our models, since a  
62 higher income gives incentives to commute. In conclusion, we filter the initial flow matrix for inter-  
63 communal travel trips, income data available only in 1595 communes and all zero flows, which gives a  
64 final sample size of 46,659 OD flows. Clearly, this is a limitation of our modelling approach but  
65 nevertheless the findings can be of apparent value for pointing directions.  
66 Modelling commuting behaviour requires a set of relevant explanatory variables that capture  
67 the characteristics of origins and destinations, along with the mechanisms that generate the trips  
68 among them. The dependent variable, inter-communal travel flows, is regressed on several  
69 independent variables obtained or derived from the 2000 Federal Census, the Swiss national transport  
70 model ARE (2005), and the Institute for Transport Planning and Systems (IVT) of ETH Zurich. We use  
71 the following variables in our framework, which are also common in explaining public transport  
72 demand in the literature (e.g. LeSage and Thomas-Agnan, 2015; Farmer, 2011; Axhausen et al.,  
73 2015).

74  
75 Table 1: Model variables  
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Statistic	Definition
Flow	Av. daily flows
Network dist.	minutes
Income diff. rel	CHF (in 1,000)
Population (o)	# inhabitants
Jobs (d)	# jobs
Pop. density (d)	# pop. / area (in km <sup>2</sup> )
Job density (o)	# jobs / area (in km <sup>2</sup> )
Pop. accessibility (d)	# accessible pop.
Job accessibility (o)	# accessible jobs
Car (o)	# cars / pop.
Car (d)	# cars / pop.
Jobs3rd (o)	# jobs in 3rd sector / # jobs
Jobs3rd (d)	# jobs in 3rd sector / # jobs
Workers (o)	# workers3 / pop.
Workers (d)	# workers3 / pop.

Note: (o),(d) = at origin, at destination municipalities

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The starting point is a log form least-squares gravity model for OD flows in the form of

$$80 \quad \log(y) = \alpha \log(l_N) + \beta_o \log(X_o) + \beta_d \log(X_d) + \delta \left( \frac{inc_d - inc_o}{inc_o} \right) + \gamma \log(g) + \epsilon,$$

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82 where  $X_o$  and  $X_d$  are characteristics of origins and destinations,  $g$  denotes the network distance and  
83  $((inc_d - inc_o)/inc_o)$  reflects the relative difference of income between destination and origin  
84 municipalities. Estimation results of the gravity model are showed in Table 2. All parameters are highly  
85 significant except those of the share of 3rd sector jobs at origins and the share of cars per origin  
86 municipality having p-values lower than 5% and 1% respectively. The network distance decay  
87 parameter (-1.537) is within the expected range for commuting patterns, in accordance with previous  
88 studies. All other explanatory variables have a much weaker impact on the dependent variable, but  
89 this finding is in line with the expectations of existing literature (LeSage and Thomas-Agnan, 2015;  
90 Farmer, 2011). Income differences between destinations and origins have a significant and positive  
91 effect on travel-to-work trips and should be interpreted as an elasticity, since relative differences are  
92 used. This intuitively makes sense. By applying Moran's I tests we find that the residuals of the  
93 aspatial model indeed exhibit remaining spatial dependence and thus justify the need for spatial  
94 models. Spatial autoregressive models (SAR) are typically written as

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$$96 \quad y = \alpha l_N + \rho_i W_i y + \beta_o X_o + \beta_d X_d + \epsilon, \quad \text{with } i = o, d.$$

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98 where the weights for  $W_i$  are defined as:

100 *Netw. dist. weights:*  $w_{ij} = \frac{1}{traveltime_{ij}}$ , *Econ. dist. weights:*  $w_{ij} = \left( \frac{traveltime_{ij}}{\exp(((inc_d - inc_o)/inc_o))} \right)^{-1}$

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102 As it can be seen in Table 2, SAR models relying on origin- and destination-centric network and  
 103 economic distance weights show positive influence of neighbouring communes on travel-to-work trips.  
 104 Rho is higher than 1, which is an artefact of using the min-max approach for the spatial weights when  
 105 building  $W_i$ ,  $i = (o,d)$  instead of classic row-normalization (Kelejian and Prucha, 2010). In the transition  
 106 from the gravity model to the SAR models, variables Car (o), Car (d), and Jobs3rd (o) are not  
 107 statistically significant anymore and the impact of network distance becomes smaller. Interestingly, rho  
 108 for the SAR model relying on economic distance weights has a bigger impact compared to the network  
 109 distance weighted SAR. It has to be emphasized that parameter estimates of spatial autoregressive  
 110 regression models can not be interpreted as simple elasticities as in the gravity model, since spatial  
 111 spillovers complicate the task of interpreting estimates from these models in a direct way.  
 112 Furthermore, the spatial models yield a higher goodness-of-fit measure than the gravity model. Finally,  
 113 we employ an IV regression framework to test the endogeneity of income and using a set of valid  
 114 instruments (in line with Sarlas and Axhausen, 2017). Note that pseudo  $R^2$  values must be treated with  
 115 caution, as they are not equivalent to OLS-based  $R^2$  measures.

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117 Table 2: Gravity model and spatial autoregressive models estimates

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*Dependent variable: log(commuting flows)*

	<b>Gravity model (OLS)</b>		<b>SAR (GMM / 2IV)</b>		<b>SAR (GMM / 2IV)</b>	
	Estimate	Sign.	Estimate	Sign.	Estimate	Sign.
(Intercept)	4.443	***	5.765	***	5.788	***
log(Netw. distance)	-1.537	***	-1.250	***	-1.254	***
Rel. Income diff.	0.085	***	0.047	**	0.041	**
log(Jobs) (d)	0.473	***	0.307	***	0.308	***
log(Pop. density) (d)	0.030	***	0.036	***	0.036	***
log(Pop. access.) (d)	-0.176	***	-0.207	***	-0.206	***
log(Jobs3rd) (d)	0.102	***	0.082	***	0.082	***
log(Car) (d)	-0.071	***	-0.007		-0.006	
log(Workers) (d)	0.665	***	0.409	***	0.417	***
log(Population) (o)	0.440	***	0.359	***	0.358	***
log(Job density) (o)	-0.043	***	-0.042	***	-0.042	***
log(Job access.) (o)	-0.180	***	-0.239	***	-0.239	***
log(Jobs3rd) (o)	-0.027	**	-0.019		-0.018	
log(Car) (o)	-0.023	*	-0.003		-0.002	
log(Workers) (o)	0.365	***	0.249	***	0.254	***
rho			2.387	***	2.704	***
$R^2$	0.5177					
Pseudo adj. $R^2$			0.5898		0.5893	
HC robust std. errors	yes		yes		yes	

Note:

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120 The resulting in-sample predictive accuracies outperform those from the current NPVM for different  
 121 accuracy measures, as can be seen in Table 3.

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123 Table 3: In-sample predictive accuracy measures

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	RMSPE	RMdSPE	MAPE	MdAPE	SMAPE	SMdAPE
NPVM	87.50	6.49	271.76	64.92	74.12	64.24
Gravity model (OLS)	131.20	68.76	913.60	687.62	143.36	154.94
SAR (GMM / 2IV); Netw. dist. weights	6.81	4.75	51.53	47.52	62.59	55.22
SAR (GMM / 2IV); Econ. dist. weights	6.40	5.13	51.99	51.34	68.96	63.03

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