Testing -6.3 *t*-tests

Michel Bierlaire

Solution to the practice quiz

The estimation results are summarized in Table 1.

| | | | Robust | | |
|-----------|---|----------|------------|--------|---------|
| Parameter | | Coeff. | Asympt. | | |
| number | Description | estimate | std. error | t-stat | p-value |
| 1 | One stop–same airline dummy | -1.17 | 0.278 | -4.19 | 0.00 |
| 2 | One stop–multiple airlines dummy | -1.45 | 0.292 | -4.98 | 0.00 |
| 3 | Elapsed time (hours) (non stop) | -0.341 | 0.0854 | -3.99 | 0.00 |
| 4 | Elapsed time (hours) (one stop-same airline) | -0.291 | 0.0822 | -3.54 | 0.00 |
| 5 | Elapsed time (hours) (one stop–multiple airlines) | -0.310 | 0.0802 | -3.87 | 0.00 |
| 6 | Round trip fare (\$100) | -1.78 | 0.151 | -11.84 | 0.00 |
| 7 | Leg room (inches), if male | 0.108 | 0.0232 | 4.65 | 0.00 |
| 8 | Leg room (inches), if female | 0.132 | 0.0221 | 5.99 | 0.00 |
| 9 | Being early (hours) | -0.151 | 0.0188 | -8.02 | 0.00 |
| 10 | Being late (hours) | -0.0960 | 0.0167 | -5.73 | 0.00 |
| 11 | More than 2 air trips per year (one stop–same airline) | -0.307 | 0.141 | -2.18 | 0.03 |
| 12 | More than 2 air trips per year (one stop–multiple airlines) | -0.0910 | 0.157 | -0.58 | 0.56 |
| 13 | Male dummy (one stop–same airline) | 0.199 | 0.126 | 1.59 | 0.11 |
| 14 | Male dummy (one stop–multiple airlines) | 0.293 | 0.132 | 2.21 | 0.03 |
| 15 | Round trip fare / income $(\$100/\$1000)$ | -24.0 | 8.09 | -2.97 | 0.00 |
| Summary | statistics | | | | |

Number of observations = 2544 $\mathcal{L}(0) = -2794.870$ $\begin{array}{c} \mathcal{L}(0) \\ \mathcal{L}(c) \\ \mathcal{L}(\widehat{\beta}) \\ -2[\mathcal{L}(0) - \mathcal{L}(\widehat{\beta})] \\ \rho^2 \\ \overline{\rho}^2 \end{array}$ -2203.160= = -1641.932= 2305.875 = 0.413 0.407 =

Table 1: Estimation results

1. Testing the null hypothesis that the true value of the coefficient of the variable "being early" is zero requires a t-test. The t statistic of parameter number 9 in Table 1 is -8.02 which is larger in absolute value than 2.56, so the null hypothesis can be rejected at the 1% level. Actually, the fact that the p value is so small that the two first digits after the decimal point are zero, is a sign that the hypothesis can be safely rejected at any reasonable level. The variable plays a role in the model.

2. The next three questions require a *t*-test to compare two coefficients β_i and β_j . The null hypothesis is that both parameters are equal $(H_0 : \beta_i = \beta_j)$ and the *t*-statistic is given by

$$\frac{\widehat{\beta}_i - \widehat{\beta}_j}{\sqrt{\operatorname{Var}(\widehat{\beta}_i - \widehat{\beta}_j)}}$$

where

$$\operatorname{Var}(\widehat{\beta}_i - \widehat{\beta}_j) = \operatorname{Var}(\widehat{\beta}_i) + \operatorname{Var}(\widehat{\beta}_j) - 2\operatorname{Cov}(\widehat{\beta}_i, \widehat{\beta}_j).$$

The variance of a parameter is the square of its standard error. The complete variance-covariance matrix can be found in v634_Boeing_MO.html. It is reported in Table 2 for the involved coefficients.

| | β_3 | β_4 | β_5 |
|-----------|-----------|-----------|-----------|
| β_3 | 0.00729 | 0.00627 | 0.006 |
| β_4 | 0.00627 | 0.00676 | 0.00553 |
| | 0.006 | | 0.00643 |

Table 2: Variance covariance matrix for the involved coefficients

The three *t*-tests are applied below. $H_0: \beta_3 = \beta_4$

$$\frac{\widehat{\beta}_3 - \widehat{\beta}_4}{\sqrt{\operatorname{Var}(\widehat{\beta}_3 - \widehat{\beta}_4)}} = \frac{-0.341 - (-0.291)}{\sqrt{0.00729 + 0.00676 - 2 \times 0.00627}} = -1.287,$$

and the *p*-value is 0.2. The null hypothesis can be rejected only at the 20% level. It is therefore reasonable not to reject it.

3. $H_0: \beta_4 = \beta_5$

$$\frac{\widehat{\beta}_4 - \widehat{\beta}_5}{\sqrt{\operatorname{Var}(\widehat{\beta}_4 - \widehat{\beta}_5)}} = \frac{-0.291 - (-0.310)}{\sqrt{0.00676 + 0.00643 - 2 \times 0.00553}} = 0.412,$$

and the *p*-value is 0.68. The null hypothesis can be rejected only at the 68% level. It is therefore reasonable not to reject it.

4.
$$H_0: \beta_3 = \beta_5$$

$$\frac{\widehat{\beta}_3 - \widehat{\beta}_5}{\sqrt{\operatorname{Var}(\widehat{\beta}_3 - \widehat{\beta}_5)}} = \frac{-0.341 - (-0.310)}{\sqrt{0.00729 + 0.00643 - 2 \times 0.006}} = -0.747,$$

and the *p*-value is 0.46. The null hypothesis can be rejected only at the 46% level. It is therefore reasonable not to reject it.

In conclusion, we have no evidence from the data that suggests that the elapsed time is not generic. Consequently, in such a circumstances, it may be worth investigating a model with a generic elapsed time, that will be more parsimonious.